

Matrix characterization of Asymptotically Probability Equivalent for Measurable Functions

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Abstract In this article, we put forward new ideas of asymptotically equivalent functions in probability and asymptotic regular matrices in probability. Two nonnegative real-valued function $f(s)$ and $g(s)$, measurable on $(1, \infty)$ are said to be asymptotically probability equivalent if for every $\varepsilon > 0$,

$$\lim_s P \left(\left| \frac{f(s)}{g(s)} - 1 \right| < \varepsilon \right) = 1$$

Also, we examine bivariate function transformation of asymptotically equivalent in probability for measurable functions.

1 Introduction

The definitions of asymptotically equivalent sequences and asymptotic regular matrices which preserve the asymptotic equivalence of two nonnegative number sequences were presented by Pobyvanets [3] and new ways of comparing rates of convergence was introduced by Fridy [1]. Furthermore, the definitions for asymptotically statistical equivalent sequences and asymptotically statistical regular matrices was presented by Patterson [4] and this concept was deal with probability theory by Patterson and Savas [5].

In order to present the main results, we first consider some definitions.

Definition 1.1. [1] For $y = (y_r)$ in \mathcal{I} , where $\mathcal{I} = \left\{ y_r : \sum_{r=1}^{\infty} |y_r| < \infty \right\}$, the "remainder sequence" $[Ry]$ is defined to be $(R_m y)$ where, for each $m \in \mathbb{N}$,

$$R_m y := \sum_m^{\infty} |y_r|.$$

Definition 1.2. [5] Two nonnegative sequences $y = (y_r)$ and $z = (z_r)$ are said to be asymptotically equivalent in probability if

$$\lim_r P \left(\left| \frac{y_r}{z_r} - 1 \right| < \varepsilon \right) = 1.$$

(denoted by $y \overset{\text{probability}}{\sim} z$)

The main goal of this paper is to introduce the ideas of asymptotically probability equivalent and asymptotically regular in probability by using $g(s)$ and $h(s)$ two nonnegative real-valued Lebesgue measurable functions on $(1, \infty)$ instead of sequences. In addition, we shall also prove some of inclusion theorems.

2 Main Definitions

In this section, we will present useful main definitions. Let us note the following notations:

$$L^f = \left\{ g(s) : \int_1^\infty |g(s)| ds < \infty \right\},$$

$$D_A = \left\{ g(s) : \lim_t \int_1^\infty \tau(t, s)g(s)ds \text{ exists, where } \tau(t, s) \text{ is a bivariate function} \right\},$$

$P_\delta^f = \{ \text{The class of all real valued lebesgue measurable functions such that } g(s) \geq \delta > 0 \text{ for all } s \},$
and

$P_0^f = \{ \text{The class of all real valued measurable functions which have at most a finite number of zero values} \}.$

Definition 2.1. Let $g(s)$ be a real-valued function which is measurable on $(1, \infty)$. If $g(s)$ is a L^f , then we denote the remainder function given by

$$R_t(g) := \int_t^\infty |g(s)| ds.$$

Definition 2.2. Let $g(s)$ be a real-valued function which is measurable on $(1, \infty)$, if $R_t(\tau(t, s)g(s))$ is a L^f , then we denote the remainder function given by

$$R_t(\tau(t, s)g(s)) := \int_t^\infty |\varphi g(s)| ds$$

when $(\varphi g)_t = \int_1^\infty \tau(t, s)g(s)ds$, in short $(\varphi g)_t$.

Definition 2.3. Two nonnegative real-valued function $g(s)$ and $h(s)$, measurable on $(1, \infty)$ are said to be asymptotically equivalent if

$$\lim_s \frac{g(s)}{h(s)} = 1$$

(denoted by $g \stackrel{F}{\sim} h$).

Definition 2.4. Two nonnegative real-valued function $g(s)$ and $h(s)$, measurable on $(1, \infty)$ are said to be asymptotically probability equivalent of multiple ξ provided that for every $\varepsilon > 0$,

$$\lim_s P \left(\left| \frac{g(s)}{h(s)} - \xi \right| < \varepsilon \right) = 1$$

(denoted by $g \stackrel{probability}{\sim} h$), and simply asymptotically probability equivalent if $\xi = 1$.

Definition 2.5. A bivariate function $\varphi = \tau(t, s)$ is asymptotically regular in probability provided that $\varphi g \stackrel{probability}{\sim} \varphi h$ whenever $g \stackrel{probability}{\sim} h$, $g(s) \in P_0^f$, and $h(s) \in P_\delta^f$ for $\delta > 0$.

3 Main Results

In the following theorem, we present necessary and sufficient conditions on the entries of a summability bivariate functions to ensure that the bivariate function transformation will conserve asymptotically equivalent in probability of multiple ξ for Lebesgue measurable functions.

Theorem 3.1. If $\varphi = \tau(t, s)$ is a bivariate function that maps bounded functions $g(s)$ into L^f , then the followings are equivalent:

- (i) If $g(s)$ and $h(s)$ are bounded functions such that $g \stackrel{probability}{\sim} h$, $g(s) \in P_0^f$, and $h(s) \in P_\delta^f$ for some $\delta > 0$, then

$$R_t(\varphi g) \stackrel{probability}{\sim} R_t(\varphi h).$$

(ii)

$$\lim_{t \rightarrow \infty} P \left(\left| \frac{\int_{m=t}^{\infty} \tau(m, p) dm}{\int_{m=t}^{\infty} \int_{n=1}^{\infty} \tau(m, n) dndm} < \varepsilon \right. \right) = 1 \text{ for each } p \text{ and } \varepsilon > 0.$$

Proof. The definition for asymptotically equivalent in probability of multiple ξ can be defined as the following

$$\lim_s P \left(\left| \frac{g(s)}{h(s)} - \xi \right| < \varepsilon \right) = 1$$

This implies that

$$\lim_s P ((\xi - \varepsilon) h(m) < g(m) \leq (\xi + \varepsilon)h(m)) = 1. \tag{3.1}$$

We now consider the following

$$\frac{R_t(\varphi g)}{R_t(\varphi h)} \leq \frac{\int_1^{N-1} g(n) \int_{m=t}^{\infty} \sup_{0 \leq n \leq N-1} \tau(m, n) dmdn}{\int_t^{\infty} \int_1^{\infty} \tau(m, n) h(n) dndm} + \frac{\int_t^{\infty} \int_N^{\infty} \tau(m, n) g(n) dndm}{\int_t^{\infty} \int_1^{\infty} \tau(m, n) h(n) dndm}$$

and we consider lower bound as follows:

$$\begin{aligned} R_t(\varphi g) &= \int_t^{\infty} \int_1^{\infty} \tau(m, n) dndm \\ &\leq \int_{m=t}^{\infty} \inf_{0 \leq n \leq N-1} \tau(m, n) dmdn \int_1^{N-1} g(n) + \int_{m=t}^{\infty} \int_N^{\infty} \tau(m, n) dndm. \end{aligned}$$

From (3.1) we are granted the following

$$\frac{R_t(\varphi g)}{R_t(\varphi h)} \leq \frac{\int_{n=1}^{N-1} g(n) \int_{m=t}^{\infty} \sup_{0 \leq n \leq N-1} \tau(m, n) dmdn}{\delta \int_{m=t}^{\infty} \int_1^{\infty} \tau(m, n) dndm} + (\xi + \varepsilon) \text{ in probability}$$

and

$$\begin{aligned} \frac{R_t(\varphi g)}{R_t(\varphi h)} &\geq \left(\frac{\int_{n=1}^{N-1} g(n) dn}{\sup h(n)} \right) \left(\frac{\int_{m=t}^{\infty} \left(\inf_{0 \leq n \leq N-1} \tau(m, n) \right) dm}{\int_{m=t}^{\infty} \int_{n=1}^{\infty} \tau(m, n) dndm} \right) + (\xi + \varepsilon) \\ &\quad - \left(\frac{(\xi - \varepsilon) \int_1^{N-1} h(n) dn}{\delta} \right) \left(\frac{\int_{m=t}^{\infty} \left(\max_{0 \leq n \leq N-1} \tau(m, n) \right) dm}{\int_{m=t}^{\infty} \int_1^{\infty} \tau(m, n) dndm} \right) \text{ in probability.} \end{aligned}$$

Therefore, from the last two equations we get the following

$$\lim_t P \left(\left| \frac{R_t(\varphi g)}{R_t(\varphi h)} - \xi \right| < \varepsilon \right) = 1$$

Hence,

$$R_t(\varphi g) \overset{\text{probability}}{\sim} R_t(\varphi h).$$

For the second part of this theorem, we consider the following two functions,

$$g(m) := \begin{cases} 0, & m \leq p; \\ 1, & \text{otherwise} \end{cases}$$

where $p \in \mathbb{Z}^+$ and $h(m) = 1$ for all m . These two functions imply as follows:

$$\begin{aligned} R_t(\varphi g) &= \int_{m=t}^{\infty} \int_{n=p+1}^{\infty} \tau(m, n) dndm \\ &= \int_{m=t}^{\infty} \int_{n=1}^{\infty} \tau(m, n) dndm - \int_{m=t}^{\infty} \int_{n=1}^p \tau(m, n) dndm. \end{aligned}$$

Therefore

$$\liminf_t \frac{R_t(\varphi g)}{R_t(\varphi h)} \leq 1 - \limsup_{t \rightarrow \infty} \frac{\int_{m=t}^{\infty} \tau(m, p) dm}{\int_{m=t}^{\infty} \int_{n=1}^{\infty} \tau(m, n) dndm}$$

Since each nonconstant element of the last inequality has statistically limit zero, We have the following

$$\lim_t P \left(\left| \frac{R_t(\varphi g)}{R_t(\varphi h)} - 1 \right| < \varepsilon \right) = 0.$$

□

Theorem 3.2. *In order for bivariate function $A = \tau(t, s)$ to be asymptotically probability regular it is necessary and sufficient that for each fixed positive integer s_0*

1. $\int_1^{s_0} \tau(m, t) dm$ is bounded for each m .
- 2.

$$\lim_t P \left(\left| \frac{\int_{t=1}^{s_0} \tau(m, t) dm}{\int_1^{\infty} \tau(m, t) dm} < \varepsilon \right| \right) = 1$$

Proof. The necessary part of above theorem is similarly established as necessary part of the previous theorem. To construct the sufficient part, let $\varepsilon > 0$, $g \stackrel{\text{probability}}{\sim} h$, $g(s) \in P_0^f$ and $h(s) \in P_\delta^f$ for some $\delta > 0$, we can write

$$P((\xi - \varepsilon)h(m + \alpha) \leq g(m + \alpha) \leq (\xi + \varepsilon)h(m + \alpha)) = 1. \tag{3.2}$$

Let us consider the following

$$\begin{aligned} \frac{(\varphi g)_t}{(\varphi h)_t} &= \frac{\int_{t=1}^{\alpha} \tau(m, t)g(t)dt + \int_{t=\alpha+1}^{\infty} \tau(m, t)g(t)dt}{\int_{t=1}^{\alpha} \tau(m, t)h(t)dt + \int_{t=\alpha+1}^{\infty} \tau(m, t)h(t)dt} \\ &= \frac{\int_{t=1}^{\alpha} \tau(m, t)g(t)dt}{\int_{t=\alpha+1}^{\infty} \tau(m, t)h(t)dt} + \frac{\int_{t=\alpha+1}^{\infty} \tau(m, t)g(t)dt}{\int_{t=\alpha+1}^{\infty} \tau(m, t)h(t)dt} \\ &= \frac{\int_{t=1}^{\alpha} \tau(m, t)h(t)dt}{\int_{t=\alpha+1}^{\infty} \tau(m, t)h(t)dt} + 1 \end{aligned}$$

Inequality (3.2) implies that

$$\lim_{m=t} P \left(\left| \frac{\int_{t=\alpha+1}^{\infty} \tau(m, t) g(t) dt}{\int_{t=\alpha+1}^{\infty} \tau(m, t) h(t) dt} - \xi \right| < \varepsilon \right) = 1.$$

Since $g(s) \in P_0^f$, $h(s) \in P_\delta^f$, and condition (2) holds we obtain the following

$$\lim_t P \left(\left| \frac{\int_{t=1}^{\alpha} \tau(m, t) g(t) dt}{\int_{t=\alpha+1}^{\infty} \tau(m, t) h(t) dt} \right| < \varepsilon \right) = 1,$$

and

$$\lim_t P \left(\left| \frac{\int_{t=1}^{\infty} \tau(m, t) h(t) dt}{\int_{t=\alpha+1}^{\infty} \tau(m, t) h(t) dt} \right| < \varepsilon \right) = 1.$$

Thus

$$\lim_t P \left(\left| \frac{(\varphi g)_t}{(\varphi h)_t} - \xi \right| < \varepsilon \right) = 1.$$

This implies that $(\varphi g) \overset{\text{probability}}{\sim}_t (\varphi h)$ whenever $g \overset{\text{probability}}{\sim} h$, $g \in P_0^f$ and $h \in P_\delta^f$ for some $\delta > 0$. \square

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